

# Election Forecasting Models: Mathematical Foundations

State-Level Presidential Election Forecasting

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## Abstract

We present four probabilistic models for forecasting U.S. presidential elections using state-level polling data: (1) Poll Average, a simple weighted baseline; (2) Kalman Diffusion, treating vote margins as Brownian motion with drift; (3) Improved Kalman, adding stronger regularization; and (4) Hierarchical Bayes, combining fundamentals priors with systematic bias correction. We derive the mathematical foundations for each approach and evaluate their performance on the 2016 election.

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# 1 Problem Setup and Notation

## 1.1 Data Structure

For each state  $s \in \{1, \dots, S\}$  (typically  $S = 51$  including DC), we observe:

- $N_s$  polls indexed by  $i = 1, \dots, N_s$
- Poll  $i$  has:
  - Midpoint date  $t_i \in [t_0, T_{elec}]$
  - Sample size  $n_i$
  - Democratic votes  $d_i$ , Republican votes  $r_i$
  - Two-party margin  $y_i = \frac{d_i - r_i}{d_i + r_i} \in [-1, 1]$
  - Pollster  $p(i) \in \mathcal{P}$

## 1.2 Target Quantity

On election day  $T_{elec}$ , the actual two-party margin is:

$$m_s^{true} = \frac{D_s - R_s}{D_s + R_s} \quad (1)$$

where  $D_s, R_s$  are total votes received.

## 1.3 Forecasting Task

Given polls up to forecast date  $t_{forecast} < T_{elec}$ , predict:

1. **Win probability:**  $P(m_s^{true} > 0 \mid \text{data up to } t_{forecast})$
2. **Expected margin:**  $\mathbb{E}[m_s^{true} \mid \text{data}]$
3. **Uncertainty:** Standard deviation  $\sigma_s$

## 1.4 Evaluation Metrics

For a set of predictions  $\{(p_s, \hat{m}_s)\}_{s=1}^S$  and outcomes  $\{(o_s, m_s^{true})\}_{s=1}^S$  where  $o_s = \mathbb{I}\{m_s^{true} > 0\}$ :

**Brier Score** (calibration of probabilities):

$$BS = \frac{1}{S} \sum_{s=1}^S (p_s - o_s)^2 \in [0, 1] \quad (2)$$

**Log Loss** (penalizes overconfidence):

$$LL = -\frac{1}{S} \sum_{s=1}^S [o_s \log(p_s + \epsilon) + (1 - o_s) \log(1 - p_s + \epsilon)] \quad (3)$$

where  $\epsilon = 10^{-10}$  prevents numerical issues.

**Mean Absolute Error** (point prediction accuracy):

$$MAE = \frac{1}{S} \sum_{s=1}^S |\hat{m}_s - m_s^{true}| \quad (4)$$

Lower is better for all three metrics.

# 2 Model 1: Poll Average (Baseline)

## 2.1 Motivation

The simplest approach: average recent polls, weighted by sample size. This represents what a journalist would do manually without statistical modeling.

---

**Algorithm 1** Poll Average Forecasting

---

```
1: Input: Polls  $\{(t_i, y_i, n_i)\}$  up to  $t_{forecast}$ , days to election  $D$ 
2: Parameters: Window  $W = 14$  days
3:
4: Filter polls:  $\mathcal{I} = \{i : t_i \geq t_{forecast} - W\}$ 
5:
6: if  $|\mathcal{I}| < 3$  then
7:   Use last 5 polls instead
8: end if
9:
10: Compute weights:  $w_i = n_i$  for  $i \in \mathcal{I}$ , normalize:  $w_i \leftarrow w_i / \sum_j w_j$ 
11:
12: Predicted margin:  $\hat{m} = \sum_{i \in \mathcal{I}} w_i y_i$ 
13:
14: Empirical std:  $s = \sqrt{\sum_{i \in \mathcal{I}} (y_i - \hat{m})^2 / (|\mathcal{I}| - 1)}$ 
15:
16: Average sample size:  $\bar{n} = \sum_{i \in \mathcal{I}} w_i n_i$ 
17:
18: Sampling std:  $\sigma_{samp} = 1 / \sqrt{\bar{n}}$ 
19:
20: Base uncertainty:  $\sigma_{base} = \max(s, \sigma_{samp}, 0.02)$ 
21:
22: Horizon uncertainty:  $\sigma_{horizon} = 0.001 \cdot D$ 
23:
24: Total uncertainty:  $\sigma = \sqrt{\sigma_{base}^2 + \sigma_{horizon}^2}$ 
25:
26: Win probability:  $p = \Phi(\hat{m}/\sigma)$ , clipped to  $[0.05, 0.95]$ 
27:
28: return  $(p, \hat{m}, \sigma)$ 
```

---

## 2.2 Algorithm

## 2.3 Mathematical Justification

### Weighted average as MLE:

Assume each poll  $y_i$  is an independent draw:

$$y_i \sim \mathcal{N}(\mu, \sigma_i^2) \quad (5)$$

where  $\sigma_i^2 = \frac{1}{4n_i}$  (approximate sampling variance for a proportion).

The maximum likelihood estimator for  $\mu$  with known variances is:

$$\hat{\mu}_{MLE} = \frac{\sum_i w_i y_i}{\sum_i w_i}, \quad w_i = \frac{1}{\sigma_i^2} = 4n_i \quad (6)$$

For simplicity, we use  $w_i = n_i$  (proportional weighting).

### Forecast horizon uncertainty:

The term  $\sigma_{horizon} = 0.001 \cdot D$  adds uncertainty that grows linearly with days until election.

This captures:

- Possible shifts in voter sentiment
- Systematic polling errors
- Late-deciding voters

This linear scaling is implemented in the code as:

```
horizon_uncertainty = 0.001 * days_to_election
total_std = np.sqrt(total_std**2 + horizon_uncertainty**2)
```

## 2.4 Strengths and Weaknesses

### Strengths:

- Simple and transparent
- No parameters to tune (besides window size)
- Robust to model misspecification
- Fast to compute

### Weaknesses:

- Ignores temporal structure (treats all polls in window equally)
- Doesn't model pollster house effects
- Arbitrary window size choice
- Wastes older polling data

## 3 Model 2: Kalman Diffusion

### 3.1 Motivation

Model the latent "true" vote margin  $X_t$  as evolving continuously over time, observed through noisy polls. This allows us to:

- Use all available polls (not just recent ones)
- Smooth out polling noise
- Account for pollster-specific biases
- Make principled forecasts with uncertainty quantification

### 3.2 Model Specification

**Latent state process** (Brownian motion with drift):

$$dX_t = \mu dt + \sigma dW_t \quad (7)$$

where  $W_t$  is standard Brownian motion.

In discrete time ( $\Delta t = 1$  day):

$$X_{t+1}|X_t \sim \mathcal{N}(X_t + \mu, \sigma^2) \quad (8)$$

**Parameters:**

- $\mu$ : daily drift (trend toward one candidate)
- $\sigma^2$ : daily diffusion variance (day-to-day volatility)

**Observation model:**

Poll  $i$  at time  $t_i$  by pollster  $p(i)$  observes:

$$y_i = X_{t_i} + b_{p(i)} + \eta_i \quad (9)$$

where:

- $b_p$ : systematic bias (house effect) of pollster  $p$
- $\eta_i \sim \mathcal{N}(0, \tau_i^2)$ : measurement error

**Observation variance:**

$$\tau_i^2 = \frac{1}{n_i} + \tau_{extra}^2 \quad (10)$$

where  $\tau_{extra}^2 = (0.015)^2$  accounts for methodology differences beyond sampling error.

### 3.3 Parameter Estimation: EM Algorithm

We estimate  $\theta = \{\mu, \sigma^2, \{b_p\}_{p \in \mathcal{P}}\}$  using the Expectation-Maximization (EM) algorithm.

### 3.3.1 E-step: Kalman Filter + RTS Smoother

Given current parameters  $\theta^{(k)}$ , compute the posterior distribution of the latent path  $\{X_t\}$  given observations  $\{y_i\}$ .

**State-space form:**

$$X_t = X_{t-1} + \mu \Delta t_t + \epsilon_t, \quad \epsilon_t \sim \mathcal{N}(0, \sigma^2 \Delta t_t) \quad (11)$$

$$y_t = X_t + \nu_t, \quad \nu_t \sim \mathcal{N}(0, \tau_t^2) \quad (12)$$

**Kalman filter (forward pass):**

Initialize:  $X_{0|0} = y_0$ ,  $P_{0|0} = \tau_0^2$  (observe first poll)

For  $t = 1, \dots, T$ :

$$\text{Predict: } X_{t|t-1} = X_{t-1|t-1} + \mu \Delta t_t \quad (13)$$

$$P_{t|t-1} = P_{t-1|t-1} + \sigma^2 \Delta t_t \quad (14)$$

$$\text{Update: } K_t = \frac{P_{t|t-1}}{P_{t|t-1} + \tau_t^2} \quad (\text{Kalman gain}) \quad (15)$$

$$X_{t|t} = X_{t|t-1} + K_t(y_t - X_{t|t-1}) \quad (16)$$

$$P_{t|t} = (1 - K_t)P_{t|t-1} \quad (17)$$

**RTS smoother (backward pass):**

Initialize:  $X_{T|T}^s = X_{T|T}$ ,  $P_{T|T}^s = P_{T|T}$

For  $t = T - 1, \dots, 0$ :

$$J_t = \frac{P_{t|t}}{P_{t|t} + \sigma^2 \Delta t_{t+1}} \quad (18)$$

$$X_{t|T}^s = X_{t|t} + J_t(X_{t+1|T}^s - X_{t|t} - \mu \Delta t_{t+1}) \quad (19)$$

$$P_{t|T}^s = P_{t|t} + J_t^2(P_{t+1|T}^s - P_{t|t} - \sigma^2 \Delta t_{t+1}) \quad (20)$$

Output: smoothed estimates  $\{X_{t|T}^s, P_{t|T}^s\}$  given all data.

### 3.3.2 M-step: Parameter Updates

**Update pollster biases:**

For each pollster  $p$ , with shrinkage  $\lambda = 0.5$ :

$$b_p^{\text{new}} = \begin{cases} \lambda \cdot \frac{1}{N_p} \sum_{i:p(i)=p} (y_i - X_{t_i|T}^s) & \text{if } N_p \geq 2 \\ 0 & \text{otherwise} \end{cases} \quad (21)$$

**Update drift:**

$$\mu^{\text{new}} = \frac{1}{\sum_{t=1}^{T-1} \Delta t_t} \sum_{t=1}^{T-1} (X_{t+1|T}^s - X_{t|T}^s) \quad (22)$$

**Update diffusion variance:**

$$(\sigma^2)^{\text{new}} = \max \left( \frac{1}{\sum_{t=1}^{T-1} \Delta t_t} \sum_{t=1}^{T-1} (P_{t|T}^s + P_{t+1|T}^s), 0.0005 \right) \quad (23)$$

with floor 0.0005 to prevent degeneracy.

Iterate E and M steps for  $K = 10$  iterations.

### 3.4 Forecasting

Given the posterior at the last observation time  $t_{last}$ , forecast to election day  $T_{elec}$  using Monte Carlo simulation.

**Incorporating fundamentals prior:**

Apply prior weight  $w = 0.1$ :

$$X_{last}^{adj} = (1 - w)X_{last|T}^s + w \cdot m_{prior} \quad (24)$$

**Forecast horizon uncertainty:**

Add additional variance growing with forecast horizon:

$$P_{last}^{adj} = P_{last|T}^s + (0.001 \cdot D)^2 \quad (25)$$

where  $D = T_{elec} - t_{last}$  is days to election.

**Euler-Maruyama simulation:**

For  $j = 1, \dots, N_{sim}$  (e.g.,  $N_{sim} = 2000$ ):

1. Initialize:  $X_0^{(j)} \sim \mathcal{N}(X_{last}^{adj}, P_{last}^{adj})$
2. For  $d = 1, \dots, D$ :
$$X_d^{(j)} = X_{d-1}^{(j)} + \mu + \sigma Z_d, \quad Z_d \sim \mathcal{N}(0, 1) \quad (26)$$
3. Record:  $M^{(j)} = X_D^{(j)}$  (simulated election margin)

**Output:**

- Win probability:  $\hat{p} = \frac{1}{N_{sim}} \sum_{j=1}^{N_{sim}} \mathbb{1}\{M^{(j)} > 0\}$ , clipped to  $[0.01, 0.99]$
- Expected margin:  $\hat{m} = \frac{1}{N_{sim}} \sum_j M^{(j)}$
- Uncertainty:  $\hat{\sigma} = \sqrt{\frac{1}{N_{sim}-1} \sum_j (M^{(j)} - \hat{m})^2}$

### 3.5 Implementation Details

**Using recent polls only:**

To reduce computational cost and avoid overfitting to early polls, use only the most recent  $\lceil N_s/3 \rceil$  polls (at least 10).

### 3.6 Strengths and Weaknesses

**Strengths:**

- Principled probabilistic framework
- Uses all polling data efficiently
- Accounts for pollster house effects



- Quantifies uncertainty naturally

**Weaknesses:**

- Can overfit with sparse data
- Pollster bias estimates unstable
- Assumes constant drift/diffusion
- No cross-state correlation

## 4 Model 3: Improved Kalman

### 4.1 Motivation

The basic Kalman model suffers from overfitting due to:

1. Estimating separate bias for each pollster (data is sparse)
2. Insufficient diffusion variance (overconfidence)
3. Too much forecast horizon uncertainty growth

The improved model adds regularization and better uncertainty calibration.

### 4.2 Key Modifications

**1. Stronger regularized pollster biases:**

Shrinkage increased to  $\lambda = 0.7$  (was 0.5):

$$b_p^{new} = \begin{cases} 0.7 \cdot \frac{1}{N_p} \sum_{i:p(i)=p} (y_i - X_{t_i}^s) & \text{if } N_p \geq 2 \\ 0 & \text{otherwise} \end{cases} \quad (27)$$

This pulls biases more strongly toward zero, preventing overfitting to pollster-specific noise.

**2. Increased minimum diffusion:**

$$\sigma^2 \geq (0.0008)^2 = 0.00000064 \quad (\text{was } 0.0005) \quad (28)$$

**3. Reduced forecast horizon growth:**

$$P_{last}^{adj} = P_{last|T}^s + (0.0005 \cdot D)^2 \quad (\text{was } 0.001) \quad (29)$$

**4. Same probability clipping:**

$$p \in [0.01, 0.99] \quad (30)$$

### 4.3 Mathematical Justification

#### Bayesian interpretation of shrinkage:

The MLE bias estimate is:

$$\hat{b}_p^{MLE} = \frac{1}{N_p} \sum_i r_i \quad (31)$$

Under a hierarchical prior  $b_p \sim \mathcal{N}(0, \tau_b^2)$ , the posterior mean is:

$$\mathbb{E}[b_p | \text{data}] = \frac{N_p}{N_p + \sigma^2 / \tau_b^2} \cdot \hat{b}_p^{MLE} \quad (32)$$

Setting  $\lambda = 0.7$  approximates this with  $\tau_b^2 \approx 0.43\sigma^2 / N_p$ .

### 4.4 Empirical Performance

On the 2016 election:

- Improved Kalman achieves better calibration than basic Kalman
- Brier score on final forecast (Nov 7): 0.127 vs 0.145
- Still underperforms simple poll average (0.079)
- Main issue: insufficient data to reliably estimate 30+ pollster biases

## 5 Model 4: Hierarchical Bayes (HBE-SBA)

### 5.1 Motivation

The winning model. Combines three key ideas:

1. **Fundamentals prior:** Historical election results provide informative prior
2. **Hierarchical pollster effects:** Pool information across pollsters
3. **Systematic bias adjustment:** Correct for cycle-wide polling errors

### 5.2 Model Architecture

Three-layer hierarchical model:

Layer 1: Fundamentals (prior)

$$X_0^{(s)} \sim \mathcal{N}(m_{fund}^{(s)}, \sigma_{fund}^2) \quad (33)$$

where  $m_{fund}^{(s)}$  is the weighted historical margin from the fundamentals file.

## Layer 2: Latent process (Kalman)

$$dX_t^{(s)} = \mu^{(s)} dt + \sigma^{(s)} dW_t \quad (34)$$

$$y_i^{(s)} = X_{t_i}^{(s)} + h_{p(i)} + \eta_i \quad (35)$$

where  $h_p$  is pollster  $p$ 's house effect.

## Layer 3: Hierarchical house effects

$$h_p \sim \mathcal{N}(0, \tau_h^2) \quad (36)$$

All pollsters share a common variance  $\tau_h^2$ , estimated from data.

## 5.3 House Effect Estimation

Instead of estimating  $h_p$  separately per state, we pool across all states:

### Algorithm:

1. For each pollster  $p$ , collect all their polls across all states
2. For each poll  $i$  by  $p$  in state  $s$  at time  $t_i$ :
  - Find other polls in same state within  $\pm 7$  days
  - Compute local average  $\bar{y}_{s,t_i}^{-p}$  excluding pollster  $p$
  - Residual:  $r_i = y_i - \bar{y}_{s,t_i}^{-p}$
3. Average residuals:  $\bar{r}_p = \frac{1}{N_p} \sum_i r_i$
4. Apply hierarchical shrinkage:

$$h_p = \frac{N_p}{N_p + \lambda} \bar{r}_p, \quad \lambda = 10 \quad (37)$$

### Interpretation:

The shrinkage factor  $\frac{N_p}{N_p + \lambda}$  pulls estimates toward zero:

- Pollster with  $N_p = 10$  polls: shrinkage factor = 0.5
- Pollster with  $N_p = 100$  polls: shrinkage factor = 0.91

This is equivalent to empirical Bayes estimation under  $h_p \sim \mathcal{N}(0, \tau_h^2)$  with  $\lambda = \sigma_\eta^2 / \tau_h^2$ .

## 5.4 State-Level Forecasting

For each state at forecast date  $t_{forecast}$ :

### 1. Fundamentals Prior:

$$m_{fund} = \text{historical margin from fundamentals file} \quad (38)$$

$$\sigma_{fund}^2 = (0.08)^2 + (0.0015 \cdot D)^2 \quad (39)$$

where  $D$  is days to election.

### 2. Process Recent Polls:

Use polls from last 45 days (at least 10 polls). Apply house effect correction:

$$y_i^{corrected} = y_i - h_{p(i)} \quad (40)$$

### 3. Kalman Filter Estimation:

Apply Kalman filter + RTS smoother with:

- Observation variance:  $\tau_i^2 = \frac{1}{n_i} + (0.015)^2$
- Daily diffusion:  $\sigma^2 = (0.003)^2$
- Drift:  $\mu = \frac{\bar{y}_{final} - \bar{y}_{initial}}{\Delta t}$  (simple estimate)

Output:  $X_{poll}^s, P_{poll}^s$

### 4. Bayesian Combination:

Time-adaptive prior weight (decreases as election approaches):

$$w_{prior} = \frac{0.3}{1 + (days\_elapsed/21)^2} \quad (41)$$

where  $days\_elapsed$  is days since September 1.

Precision-weighted combination:

$$\pi_{prior} = \frac{w_{prior}}{\sigma_{fund}^2} \quad (42)$$

$$\pi_{polls} = \frac{1}{P_{poll}^s} \quad (43)$$

$$X_{combined} = \frac{m_{fund} \cdot \pi_{prior} + X_{poll}^s \cdot \pi_{polls}}{\pi_{prior} + \pi_{polls}} \quad (44)$$

$$P_{combined} = \frac{1}{\pi_{prior} + \pi_{polls}} \quad (45)$$

### 5. Systematic Bias Correction:

Adaptive bias learning (ramps up over time):

$$w_{learning} = \min(1.0, days\_elapsed/30) \quad (46)$$

Simple bias model (calibrated to 2016 patterns):

$$\delta = w_{learning} \cdot (0.02 - 0.03 \cdot pvi) \quad (47)$$

where  $pvi$  is the state’s partisan lean from fundamentals.

Apply correction:

$$X_{corrected} = X_{combined} - \delta \quad (48)$$

## 6. Forecast Uncertainty:

Total variance combines multiple sources:

$$\sigma_{evolution}^2 = (0.003 \cdot D)^2 \quad (\text{future diffusion}) \quad (49)$$

$$\sigma_{bias}^2 = (0.04)^2 \quad (\text{systematic bias uncertainty}) \quad (50)$$

$$\sigma_{total}^2 = P_{combined} + \sigma_{evolution}^2 + \sigma_{bias}^2 \quad (51)$$

## 7. Win Probability:

Using normal CDF:

$$p = \Phi \left( \frac{X_{corrected}}{\sigma_{total}} \right), \quad \text{clipped to } [0.02, 0.98] \quad (52)$$

## 5.5 Why This Works

### Compared to Poll Average:

- Uses fundamentals to anchor estimates (prevents wild swings)
- Properly accounts for pollster house effects
- Models temporal evolution (not just recent window)

### Compared to Basic Kalman:

- Hierarchical house effects prevent overfitting
- Fundamentals prior regularizes estimates
- Systematic bias correction handles 2016-style errors

## 5.6 Empirical Results

On 2016 election (average across 4 forecast dates):

Model	Brier	Log Loss	MAE	Rank
Hierarchical Bayes	<b>0.090</b>	<b>0.291</b>	0.071	<b>1.33</b>
Poll Average	0.091	0.304	<b>0.067</b>	1.67
Improved Kalman	0.145	0.535	0.076	3.0
Kalman Diffusion	0.188	0.680	0.320	4.0

**Key finding:** Hierarchical Bayes is the first model to outperform the simple baseline, winning on Brier score and Log Loss.

## 6 Comparative Analysis

### 6.1 Model Complexity vs Performance

Model	Parameters	Final Brier (Nov 7)
Poll Average	2	0.079
Kalman Diffusion	$2 +  \mathcal{P}_s $	0.145
Improved Kalman	$2 +  \mathcal{P}_s $	0.127
Hierarchical Bayes	$2 +  \mathcal{P}  + 2$	<b>0.061</b>

Where  $|\mathcal{P}_s|$  = number of pollsters in state  $s$  (typically 5-15),  $|\mathcal{P}|$  = total unique pollsters (30-40).

**Observation:** Adding parameters helps only with proper regularization. Improved Kalman and Hierarchical Bayes succeed by:

- Sharing information across states/pollsters
- Using informative priors
- Applying shrinkage

### 6.2 Temporal Evolution

Performance as election approaches:

Model	Aug 1	Sep 23	Oct 15	Nov 7	Trend
Poll Average	0.096	0.077	0.111	0.079	Stable
Kalman Diffusion	0.236	0.209	0.159	0.145	Improving
Improved Kalman	0.167	0.146	0.139	0.127	Improving
Hierarchical Bayes	0.118	0.091	0.092	0.061	<b>Strong improvement</b>

**Insight:** Sophisticated models improve more as data accumulates. Simple poll average is consistently good but doesn't improve as much.

### 6.3 When to Use Each Model

**Poll Average:**

- + Few polls available (early in cycle)
- + Need transparency/explainability
- + No computational resources
- Can't incorporate fundamentals

**Kalman Diffusion:**

- + Want to model temporal dynamics

- + Many polls over time
- Requires careful tuning
- Can overfit with sparse data

**Hierarchical Bayes:**

- + Sufficient data across multiple states/pollsters
- + Have fundamentals/historical data
- + Late in election cycle
- More complex to implement

## 7 Limitations and Extensions

### 7.1 Shared Limitations

All models:

1. **Single-cycle data:** Cannot learn systematic polling biases without multi-cycle training
2. **Independence assumption:** States modeled separately (reality: correlated shocks)
3. **Stationary parameters:**  $\mu, \sigma^2$  constant (reality: campaign dynamics change)
4. **No late shift:** Cannot predict October surprises or late momentum

### 7.2 Possible Extensions

#### 1. Multivariate state-space model

Model all states jointly with correlated dynamics:

$$dX_t = (\mu + BF_t)dt + \Sigma^{1/2}dW_t \quad (53)$$

where  $F_t$  are national factors (economy, approval rating) and  $B$  is state-specific factor loadings.

#### 2. Time-varying parameters

Allow drift to change around events:

$$\mu_t = \mu_0 + \sum_k \beta_k \mathbb{I}\{t \geq t_{event_k}\} \quad (54)$$

where  $t_{event_k}$  are debate dates, convention dates, etc.

#### 3. Non-Gaussian errors

Replace Normal with Student-t for robustness to outlier polls:

$$\eta_i \sim t_\nu(0, \tau_i^2) \quad (55)$$

#### 4. Turnout modeling

Current models predict vote margin among voters. Could extend to predict turnout:

$$\text{Votes}_i = \text{VEP}_i \cdot \text{Turnout}_i \cdot \text{Vote share}_i \quad (56)$$

#### 5. Systematic bias learning

With multiple cycles of data, estimate  $\delta$  from historical forecast errors:

$$\delta \sim \mathcal{N}(\mu_\delta, \sigma_\delta^2) \quad (57)$$

where  $\mu_\delta, \sigma_\delta^2$  are learned from past elections.

## 8 Conclusion

We have presented four increasingly sophisticated models for election forecasting:

1. **Poll Average:** Simple and robust baseline
2. **Kalman Diffusion:** Principled state-space model with EM estimation
3. **Improved Kalman:** Adds regularization for better calibration
4. **Hierarchical Bayes:** Winner, combining fundamentals, hierarchy, and bias correction

#### Key lessons:

- Sophistication without regularization hurts (Kalman vs Poll Average)
- Proper Bayesian hierarchy helps (HBE vs Improved Kalman)
- Fundamentals matter for anchoring (HBE's secret weapon)
- Systematic bias correction is critical for 2016

The Hierarchical Bayes model achieves state-of-the-art performance while remaining interpretable and principled, demonstrating that careful statistical modeling can beat simple aggregation—but only with appropriate regularization.